

Structural Estimation: EM Algorithm

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Last Time

- Non-Parametric estimation of mixing distribution
 - We discretize G
 - $\{\nu_j\}_{j=1}^K$: set of points in G
 - $\{\pi_j\}_{j=1}^K$: the probability of point j
- Sum over the points to get the full distribution of durations

$$f(t|x) = \sum_{j=1}^K \pi_j \times f(t|x, \nu_j)$$

- The likelihood function will be a function of $\{\nu_j\}_{j=1}^K$ and $\{\pi_j\}_{j=1}^K$ and we get ML estimates of each point and its probability.

Last Time

- **Problem:** can no longer log the likelihood function
- Increases computational burden
 - In our case $f(t|x, \nu_j)$ was “simple” enough
- **Solution:** Expectation-Maximization Algorithm

General Latent Variable Problem

- x is an observed random variable
- z is an unobserved (latent) variable
- The joint probability is parameterized by $\theta \in \Theta$

$$p(x, z; \theta)$$

- There are two sets of unknowns: z and θ
- EM algorithm
 - Guess z , maximize w.r.t. θ
 - Use the estimate of θ to get a better guess for z

Simple Example: Gaussian Mixture Model

- We observe data $x = \{x_1, \dots, x_n\}$ which are i.i.d. draws
 - $N(\mu_1, \sigma_1)$ with probability π
 - $N(\mu_2, \sigma_2)$ with probability $1 - \pi$
- We do not know which distribution each x_i came from
- So we need to estimate $\theta = \{\pi, \mu_1, \mu_2, \sigma_1, \sigma_2\}$

Simple Example: Gaussian Mixture Model

- The likelihood function, $\phi(\cdot)$ is the normal pdf

$$\begin{aligned}L(\theta; \mathbf{x}) &= \prod_{i=1}^N p(x_i; \theta) \\ &= \prod_{i=1}^N \pi \phi(x_i; \mu_1, \sigma_1) + (1 - \pi) \phi(x_i; \mu_2, \sigma_2)\end{aligned}$$

- The log-likelihood

$$\mathcal{L}(\theta; \mathbf{x}) = \sum_{i=1}^N \log[\pi \phi(x_i; \mu_1, \sigma_1) + (1 - \pi) \phi(x_i; \mu_2, \sigma_2)]$$

- **Problem:** can't distribute the log any further

Simple Example: Gaussian Mixture Model

- **Solution:** introduce a latent variable

$$z_i = \begin{cases} 1 & x_i \text{ is a draw from } N(\mu_1, \sigma_1) \\ 0 & x_i \text{ is a draw from } N(\mu_2, \sigma_2) \end{cases}$$

where $P(z_i = 1) = \pi$

- If we observed z_i then the likelihood is

$$L(\theta; x, z) = \prod_{i=1}^N [\pi \phi(x_i; \mu_1, \sigma_1)]^{z_i} [(1 - \pi) \phi(x_i; \mu_2, \sigma_2)]^{1-z_i}$$

$$\begin{aligned} \mathcal{L}(\theta; x, z) = \sum_{i=1}^N & z_i \log[\phi(x_i; \mu_1, \sigma_1)] + (1 - z_i) \log[\phi(x_i; \mu_2, \sigma_2)] \\ & + z_i \log(\pi) + (1 - z_i) \log(1 - \pi) \end{aligned}$$

Simple Example: Gaussian Mixture Model

- EM Algorithm

- 1 Guess initial values, $\hat{\theta}_0$

- 2 Expectations Step (E-Step)

- Given $\hat{\theta}_0$ estimate \hat{z}_i (we will see how to do this later)
- Construct $\mathcal{L}(\theta; x, \hat{z})$

- 3 Maximization Step (M-Step)

$$\hat{\theta}_1 = \operatorname{argmax} \mathcal{L}(\theta; x, \hat{z})$$

- 4 Repeat 2-3 until $|\hat{\theta}_{j+1} - \hat{\theta}_j| < \varepsilon$

EM Algorithm: Why does it work?

- For any θ guess, $\mathcal{L}(\theta; x, \hat{z})$ is a lower bound to $\mathcal{L}(\theta; x)$
- The algorithm is repeated maximization of lower bounds
- Two caveats
 - convergence is often slow
 - converges to local max (initial guess matters!)

EM Algorithm: Why does it work?

$$\begin{aligned}\mathcal{L}(\theta; x) &= \log P(x; \theta) \\ &= \log \sum_z P(x, z; \theta) \\ &= \log \sum_z P(z) \left(\frac{P(x, z; \theta)}{P(z)} \right) \\ &\geq \underbrace{\sum_z P(z) \log \left(\frac{P(x, z; \theta)}{P(z)} \right)}_{\mathcal{L}(\theta; x, z)} \quad (\text{Jensen's Inequality})\end{aligned}$$

- So $\mathcal{L}(\theta; x, z)$ is a lower bound for any choice of z

EM Algorithm: The best lower bound

$$\begin{aligned}\mathcal{L}(\theta; x, z) &= \sum_z P(z) \log \left(\frac{P(x, z; \theta)}{P(z)} \right) \\ &= \sum_z P(z) \log \left(\frac{P(z|x; \theta)P(x; \theta)}{P(z)} \right) \\ &= \sum_z P(z) \log \left(\frac{P(z|x; \theta)}{P(z)} \right) + \sum_z P(z) \log P(x; \theta) \\ &= -KL(P(z)||P(z|x; \theta)) + \mathcal{L}(\theta; x)\end{aligned}$$

- $KL(P(z)||P(z|x; \theta))$ is the Kullbeck-Leibler divergence
- $KL(P(z)||P(z|x; \theta)) = 0$ when $P(z) = P(z|x; \theta)$.

EM Algorithm

(E-Step) With $\hat{\theta}_j$ and compute the probabilities of z

$$P(z|x; \hat{\theta}_j) = \frac{P(x|z; \hat{\theta}_j)P(z|\hat{\theta}_j)}{\sum_z P(x|z; \hat{\theta}_j)P(z|\hat{\theta}_j)}$$

(M-Step) Maximize the lower-bound to get new estimate

$$\hat{\theta}_{j+1} = \operatorname{argmax} \sum_z P(z|x; \hat{\theta}_j) \log \left(\frac{P(x, z; \theta)}{P(z|x; \hat{\theta}_j)} \right)$$

$$\hat{\theta}_{j+1} = \operatorname{argmax} \sum_z P(z|x; \hat{\theta}_j) \log [P(x|z; \theta)P(z; \theta)]$$

EM Algorithm: Gaussian Mixture Model

- We observe $\{x_1, \dots, x_n\}$ that are i.i.d. draws from
 - $\phi(x_i, \mu_1, \sigma_1) \sim N(\mu_1, \sigma_1)$ with probability π
 - $\phi(x_i, \mu_2, \sigma_2) \sim N(\mu_2, \sigma_2)$ with probability $1 - \pi$
- If want to estimate $\theta = \{\pi, \mu_1, \mu_2, \sigma_1, \sigma_2\}$
- Introduce a latent variable

$$z_i = \begin{cases} 1 & x_i \text{ is a draw from } N(\mu_1, \sigma_1) \\ 0 & x_i \text{ is a draw from } N(\mu_2, \sigma_2) \end{cases}$$

- Start with an initial guess $\hat{\theta} = \{\hat{\pi}, \hat{\mu}_1, \hat{\mu}_2, \hat{\sigma}_1, \hat{\sigma}_2\}$

EM Algorithm

(E-Step) With $\hat{\theta}_j$ and compute the probabilities of z

$$P(z|x; \hat{\theta}_j) = \frac{P(x|z; \hat{\theta}_j)P(z|\hat{\theta}_j)}{\sum_z P(x|z; \hat{\theta}_j)P(z|\hat{\theta}_j)}$$

$$P(z = 1|x; \hat{\theta}) = \frac{\hat{\pi}\phi(x_i, \hat{\mu}_1, \hat{\sigma}_1)}{\hat{\pi}\phi(x_i, \hat{\mu}_1, \hat{\sigma}_1) + (1 - \hat{\pi})\phi(x_i, \hat{\mu}_2, \hat{\sigma}_2)}$$

$$P(z = 0|x; \hat{\theta}) = \frac{(1 - \hat{\pi})\phi(x_i, \hat{\mu}_2, \hat{\sigma}_2)}{\hat{\pi}\phi(x_i, \hat{\mu}_1, \hat{\sigma}_1) + (1 - \hat{\pi})\phi(x_i, \hat{\mu}_2, \hat{\sigma}_2)}$$

EM Algorithm

(M-Step) Maximize the lower-bound to get new estimate

$$\begin{aligned}\hat{\theta}_{j+1} &= \operatorname{argmax}_z \sum_z P(z|x; \hat{\theta}_j) \log[P(x|z; \theta)P(z; \theta)] \\ &= \operatorname{argmax} \left(P(z=1|x; \hat{\theta}) \log[\pi \phi(x_i, \mu_1, \sigma_1)] \right. \\ &\quad \left. + P(z=0|x; \hat{\theta}) \log[(1-\pi) \phi(x_i, \mu_2, \sigma_2)] \right)\end{aligned}$$

(Check) $|\hat{\theta}_{j+1} - \hat{\theta}_j| < \varepsilon$

Matlab Estimation

- using data5.csv
- File 1: SE5_main
 - Part 1: estimate the Gaussian mixture model
 - pick a ε as stopping criterion
- File 2: log_like_GM.m
 - inputs ?
 - outputs ?

Matlab Estimation: Part 1 Answer

- init guess = [1, 1, 0, 1, 0.5]

Parameter	Value	Estimate
μ_1	5	4.9760 (0.0255)
σ_1	1.2	1.1859 (0.0177)
μ_2	0	0.0029 (0.0081)
σ_2	1	1.0029 (0.0081)
π	0.2	0.1973 (0.0040)

EM Algorithm - Mixed Proportional Hazard Model

- From Last Time

$$f(t_i|x_i; \alpha, \beta, \nu_1) = \nu_1 \exp(x_i' \beta) \alpha t_i^{\alpha-1} e^{-\nu_1 \exp(x_i' \beta) t_i^\alpha}$$

$$f(t_i|x_i; \alpha, \beta, \nu_2) = \nu_2 \exp(x_i' \beta) \alpha t_i^{\alpha-1} e^{-\nu_2 \exp(x_i' \beta) t_i^\alpha}$$

$$f(t_i|x_i; \alpha, \beta, \nu_3) = \nu_3 \exp(x_i' \beta) \alpha t_i^{\alpha-1} e^{-\nu_3 \exp(x_i' \beta) t_i^\alpha}$$

- $\theta = \{\alpha, \beta, \{\nu_j\}, \{\pi_j\}\}$
- ν is our latent variable “z”

EM Algorithm

(E-Step) With $\hat{\theta}_j = \{\hat{\alpha}, \hat{\beta}, \{\hat{\nu}_j\}, \{\hat{\pi}_j\}\}$ and compute the probabilities of ν_k

$$P(z|x; \hat{\theta}_j) = \frac{P(x|z; \hat{\theta}_j)P(z|\hat{\theta}_j)}{\sum_z P(x|z; \hat{\theta}_j)P(z|\hat{\theta}_j)}$$

$$P(\nu_k|x; \hat{\theta}) = \frac{\hat{\pi}_k \hat{\nu}_k \exp(x'_i \hat{\beta}) \hat{\alpha} t_i^{\hat{\alpha}-1} e^{-\hat{\nu}_k \exp(x'_i \hat{\beta}) t_i^{\hat{\alpha}}}}{\sum_k \hat{\pi}_k \hat{\nu}_k \exp(x'_i \hat{\beta}) \hat{\alpha} t_i^{\hat{\alpha}-1} e^{-\hat{\nu}_k \exp(x'_i \hat{\beta}) t_i^{\hat{\alpha}}}}$$

EM Algorithm

(M-Step) Maximize the lower-bound to get new estimate

$$\hat{\theta}_{j+1} = \operatorname{argmax} \sum_z P(z|x; \hat{\theta}_j) \log[P(x|z; \theta)P(z; \theta)]$$

$$= \operatorname{argmax} \sum_k P(\nu_k|x; \hat{\theta}_j) \log[\pi_k \nu_k \exp(x'_i \beta) \alpha t_i^{\alpha-1} e^{-\nu_k \exp(x'_i \beta) t_i^\alpha}]$$

(Check) $|\hat{\theta}_{j+1} - \hat{\theta}_j| < \varepsilon$

Estimation Answer

- Estimates and Standard Errors

Parameter	Estimate	Standard Error
α	0.9675	0.0275
ν_1	0.0574	0.0076
ν_2	0.2565	0.0354
ν_2	0.7543	0.1310
π_1	0.0523	0.0899
π_2	0.6649	1.1531
π_3	0.2827	0.4891
β_{FE}	0.0629	0.0260
β_{educ2}	0.0044	0.1167
β_{educ3}	0.0277	0.1132

- Log-Likelihood Value

$$\log L = -2.9267e + 04$$